

MATH4210: Financial Mathematics Tutorial 4

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Review on Normal r.v.

$$\begin{aligned}\mathbb{E}[Y] &= \mathbb{E}\left[\frac{1}{n} \sum_{i=1}^n X_i\right] \\ &= \frac{1}{n} \mathbb{E}\left[\sum_{i=1}^n X_i\right] \\ &= \frac{1}{n} \sum_{i=1}^n \mathbb{E}[X_i] \quad (\mathbb{E} \text{ is linear})\end{aligned}$$

Question

Assume a sequence of i.i.d. r.v.s $\{X_i\}_{i=1 \dots n}$, $X_1 \sim N(\mu, \sigma^2)$. Denote by $Y := \frac{1}{n} \sum_{i=1}^n X_i$. Find $\mathbb{E}[Y]$ and $\text{Var}(Y)$

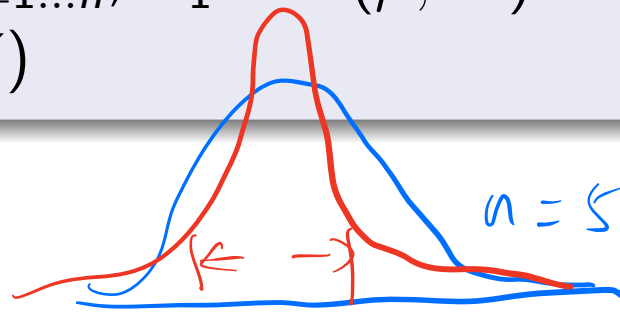
Since X_i 's are i.i.d.

So $\forall i \in [1, n]$, $\mathbb{E}[X_i] = \mu$.

Hence, $\mathbb{E}[Y] = \frac{1}{n} \cdot n \cdot \mu = \mu$.

Recall $\text{Var}\left(\sum_{i=1}^n c_i Z_i\right) = \sum_{i=1}^n c_i^2 \text{Var}(Z_i) + \sum_{1 \leq i \neq j \leq n} c_i c_j \text{Cov}(Z_i, Z_j)$

$$\text{Var}(Y) = \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i) + \frac{1}{n^2} \sum_{1 \leq i \neq j \leq n} \underbrace{\text{Cov}(X_i, X_j)}_{=0} = \frac{1}{n^2} \cdot n \cdot \sigma^2 = \frac{\sigma^2}{n}.$$



Review on Normal r.v.

$$\textcircled{1} Y \sim N(\mu_Y, \sigma_Y^2) \Leftrightarrow \mathbb{E}[e^{i\theta Y}] = \underline{e^{i\theta\mu_Y - \frac{1}{2}\theta^2\sigma_Y^2}}.$$

Fix $\theta \in \mathbb{R}$.

$$\mathbb{E}[e^{i\theta Y}] = \mathbb{E}\left[e^{i\theta \cdot \sum_{j=1}^n X_j}\right] = \mathbb{E}\left[\prod_{j=1}^n e^{i\theta X_j}\right]$$

Question

Assume a sequence of independent r.v.s $\{X_i\}_{i=1\dots n}$, for any $i \in [|1, n|]$, $X_i \sim N(\mu_i, \sigma_i^2)$. Denote by $Y := \sum_{i=1}^n X_i$. Show that Y is gaussian. Find the parameters of Y .

Since X_j 's are independent.

$$\begin{aligned} \text{Then: } \mathbb{E}\left[\prod_{j=1}^n e^{i\theta X_j}\right] &= \prod_{j=1}^n \mathbb{E}[e^{i\theta X_j}] \\ &= \prod_{j=1}^n e^{i\theta \cdot \mu_j - \frac{1}{2}\theta^2\sigma_j^2} \\ &= e^{i\theta \sum_{j=1}^n \mu_j - \frac{1}{2}\theta^2 \sum_{j=1}^n \sigma_j^2} \\ &= \underline{e^{\underbrace{i\theta \sum_{j=1}^n \mu_j}_{\mu_Y} - \frac{1}{2}\theta^2 \underbrace{\sum_{j=1}^n \sigma_j^2}_{\sigma_Y^2}}} \end{aligned}$$

$$Y \sim N\left(\sum_{j=1}^n \mu_j, \sum_{j=1}^n \sigma_j^2\right)$$

$$\begin{array}{l} X \sim N(\mu, \sigma) \\ X = \mu + \sigma \cdot Z. \\ Z \sim N(0, 1) \end{array}$$

□

Review on Normal r.v.

① X is not defined on \mathbb{R}

② Fix $x > 0$

$$\mathbb{P}[X \leq x] = \mathbb{P}[\ln X \leq \ln x]$$

Since $\ln X \sim N(\mu, \sigma^2)$

Then $\mathbb{P}[\ln X \leq \ln x]$

$$= \int_{-\infty}^{\ln x} \frac{1}{\sqrt{2\pi}\sigma} \cdot e^{-\frac{(y-\mu)^2}{2\sigma^2}} dy$$

Question

Assume X follows log-normal distribution with parameters μ, σ^2 . Find the probability density function of X .

Recall that $X \sim \text{LN}(\mu, \sigma^2)$ if

$$\ln(X) \sim N(\mu, \sigma^2)$$

$$\frac{d}{dx} \mathbb{P}[\ln X \leq \ln x] = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(\ln x - \mu)^2}{2\sigma^2}} \cdot \frac{1}{x}$$

$$\text{So } f_X(y) = \frac{1}{\sqrt{2\pi}\sigma y} \cdot e^{-\frac{(\ln y - \mu)^2}{2\sigma^2}} \quad \text{for } y > 0$$



Review on Normal r.v.

$$\textcircled{1} \mathbb{E}[X] = 0 \quad \mathbb{E}[X^2] = 1$$

$$\text{Fix } n \in \mathbb{N}^*, \quad \mathbb{E}[X^n] = \frac{1}{\sqrt{2\pi}} \int_{\mathbb{R}} x^n \cdot e^{-\frac{x^2}{2}} dx = \frac{1}{\sqrt{2\pi}} \int_{\mathbb{R}} x^{n-1} \cdot \underbrace{x \cdot e^{-\frac{x^2}{2}}}_{v'(x)} dx$$

$$\text{Derivate by } u(x) = x^{n-1}, \quad v(x) = -e^{-\frac{x^2}{2}}$$

Question

Assume $X \sim N(0, 1)$. For $n \in \mathbb{N}^*$, find $\mathbb{E}[X^n]$.

Remark

If $Y \sim N(\mu, \sigma^2)$, we can write $\mathbb{E}[Y^n]$ explicitly by considering

$$Y = \mu + \sigma Z, \quad Z \sim N(0, 1).$$

$$v'(x) = x \cdot e^{-x^2/2}$$

Then, by IBP

$$\begin{aligned} \mathbb{E}[X^n] &= \frac{1}{\sqrt{2\pi}} \left(\left[-x^{n-1} \cdot e^{-\frac{x^2}{2}} \right]_{-\infty}^{+\infty} - \int_{\mathbb{R}} (n-1)x^{n-2} \cdot (-e^{-\frac{x^2}{2}}) dx \right) \\ &= \frac{1}{\sqrt{2\pi}} \left(0 + (n-1) \cdot \int_{\mathbb{R}} x^{n-2} e^{-\frac{x^2}{2}} dx \right) = (n-1) \mathbb{E}[X^{n-2}] \end{aligned}$$

$$E[X] = 0. \quad E[X^2] = 1.$$

$$\forall n \text{ odd, } E[X^n] = 0$$

Answer:

$$E[X^n] = \begin{cases} 0 & \text{if } n=2k-1 \text{ for some } k \\ \frac{(2k)!}{2^k \cdot k!} & \text{if } n=2k \text{ for some } k. \end{cases}$$

$$E[X^{2k-1}] = 0.$$

$$E[X^{2k}] = \frac{(2k)!}{2^k k!}$$

Convergence of r.v.s

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space. X and $\{X_n\}$ are \mathbb{R} valued (sequence of) r.v.s.

↙ Convergen a.e. in real analysis

Definition (Convergence almost surely)

Denote by $X_n \rightarrow X$ a.s. (almost surely) if

$$\mathbb{P}[\{\omega \in \Omega : \lim_{n \rightarrow \infty} X_n(\omega) = X(\omega)\}] = 1$$

$= \mathbb{P}[\{\lim_{n \rightarrow \infty} X_n = X\}] = 1$

↙ Converge. in measure

Definition (Convergence in Probability)

Denote by $X_n \rightarrow X$ in probability if for any $\rho > 0$

$$\lim_{n \rightarrow \infty} \mathbb{P}[\{\omega \in \Omega : |X_n(\omega) - X(\omega)| \geq \rho\}] = 0$$

Convergence of r.v.s

$$X_n \xrightarrow{\text{a.s.}} X \text{ if } \mathbb{P}(\{\lim_{n \rightarrow \infty} X_n = X\}) = 1$$
$$X_n \xrightarrow{\mathbb{P}} X \text{ if } \forall \rho > 0, \lim_{n \rightarrow \infty} \mathbb{P}(\{|X_n - X| \geq \rho\}) = 0$$

Question

Let X and $\{X_n\}$ be \mathbb{R} valued (sequence of) r.v.s. Assume $X_n \rightarrow X$ a.s.. Show that $X_n \rightarrow X$ in probability.

Proof: Since $X_n \rightarrow X$ a.s.

$$\text{we have } \mathbb{P}(\{\lim_{n \rightarrow \infty} X_n = X\}) = 1$$

Fix $\rho > 0$ and $n \in \mathbb{N}$.

$$A_n = \bigcup_{m \geq n} \{|X_m - X| \geq \rho\}$$

$$\{|X_n - X| \geq \rho\} \subseteq A_n.$$

$$\mathbb{P}(|X_n - x| \geq \rho) \leq \mathbb{P}(A_n). \quad \text{--- } (*)$$

If we take limit equal to 0, then ---

So we study $\mathbb{P}(A_n)$.

$\{A_n\}_{n \geq 1}$ is a decreasing sequence.

$$\mathbb{P}(A_1) \leq \mathbb{P}(\Omega) = 1.$$

$$\text{Then } \mathbb{P}\left(\bigcap_{n=1}^{+\infty} A_n\right) = \lim_{n \rightarrow +\infty} \mathbb{P}(A_n) \quad \text{--- } (**)$$

$$\text{Consider } B = \left\{ \lim_{n \rightarrow +\infty} |X_n - x| = 0 \right\}. \quad \mathbb{P}(B) = 1$$

Fix $\omega \in B$.

Since $X_n \rightarrow x$ a.s., there exists $N \geq 0$
 $|X_n(\omega) - x(\omega)| < \varepsilon$ for all $n \geq N$

$$\text{But } \forall n \geq N, \omega \notin A_n \Rightarrow B \cap \left(\bigcap_{n=1}^{+\infty} A_n\right) = \emptyset$$

$$\text{So } \mathbb{P}\left(B \cup \left(\bigcap_{n=1}^{+\infty} A_n\right)\right) \leq \mathbb{P}(\Omega) = 1$$

$$\begin{aligned} & \parallel \\ & \mathbb{P}(B) + \mathbb{P}\left(\bigcap_{n=1}^{+\infty} A_n\right) \\ & \Rightarrow \mathbb{P}\left(\bigcap_{n=1}^{+\infty} A_n\right) + 1 \leq 1 \end{aligned}$$

$$\Rightarrow \mathbb{P}\left(\bigcap_{n=1}^{+\infty} A_n\right) = 0 \quad \text{--- } (***)$$

By () () ()

$$\lim_{n \rightarrow +\infty} \mathbb{P}(|X_n - x| \geq \rho) \leq \lim_{n \rightarrow +\infty} \mathbb{P}(A_n)$$

$$\begin{aligned} & \parallel \\ & \mathbb{P}\left(\bigcap_{n=1}^{+\infty} A_n\right) \end{aligned}$$

So we conclude: $X_n \rightarrow x$ in probability.